RESEARCH Open Access



The effect of long-term care insurance on women's health outcomes in China: evidence from a disability inclusion action

Wenjing Jiang^{1,2} and Hongyan Yang^{1,2*}

Abstract

Background Long-term care insurance (LTCI) in China provides financial and care security for persons with disabilities and includes caregivers in the paid labour workforce. However, it is unclear how the LTCI affects health outcomes in female recipients, female caregivers, and female non-recipients and female non-caregivers.

Methods Using the China Health and Retirement Longitudinal Study data and staggered difference-in-differences method, we evaluated the effect of LTCI on health outcomes in women with different roles, including female recipients, female caregivers, and female non-recipients and female non-caregivers, and discussed the heterogeneity of the effect on women's health outcomes.

Results LTCI statistically significant increased self-rated health and reduced depression in women and improved the health in women with different roles by increasing self-rated health in female recipients, reducing chronic diseases in female caregivers, and reducing depression in female non-recipients and female non-caregivers. There was a more pronounced improvement in health outcomes among women in the west and northeast and women in rural village.

Conclusions After the implementation of LTCI, health outcomes in female recipients, female caregivers, and female non-recipients and female non-caregivers were improved. LTCI' improvement on women's health outcomes was heterogeneous geographically and socially. Our findings highlight the importance of delivering differentiated health interventions for women with different roles in the implementation process of LTCI and minimizing women's health inequalities in geography and society.

Keywords Women, Health outcomes, Long-term care insurance, Caregivers, Staggered difference-in-differences

^{*}Correspondence:
Hongyan Yang
yhyhyang@163.com

¹Center for Social Security Studies, Wuhan University, Wuhan
430072, China

²School of Political Science & Public Administration, Wuhan University,
Wuhan 430072, China



Jiang and Yang BMC Public Health (2024) 24:2262 Page 2 of 12

Introduction

Gender inequalities women faced put their health at risk. Women are more likely to face greater barriers in accessing health services than men. These greater barriers include, but are not limited to, restrictions on women's decision-making power and mobility, discriminatory attitudes towards women, as well as a lack of awareness on women's health needs. Therefore, women are at greater health risks and suffer from less health treatment or supportive assistance. Estimates from the World Health Organization (WHO) show that 18% of women have disability compared to only 14.2% of men [1]. Persons with disabilities suffer from ableism and stigmatization in all aspects of life, contributing to their poorer physical and mental health. In addition, due to the limitations on daily functions than other persons, persons with disabilities often rely on caregivers for daily life.

Gender inequalities affect not only the health outcomes of women with disabilities, but also the health outcomes of female caregiver. According to a joint report by the International Labour Organization (ILO) and WHO The gender pay gap in the health and care sector, women in the health and care sector face a larger gender pay gap compared to other economic sectors, earning an average of 24% less than male counterparts [2]. The report published by WHO Fair share for health and care: gender and the undervaluation of health and care work illustrates that chronic underinvestment in health and care is exacerbating the global care crisis, women undertake 67% of the global paid health and care workforce, and 76% of the global unpaid care activities, which reduces women's participation in the paid labour market and hinders gender equality, and gender inequality in health and care work negatively affects women and health outcomes [3]. Compared with offspring caregivers, spousal caregivers are more likely to live with their recipients [4], experience more social isolation [5], higher financial and physical burden, as well as depressive symptoms [6]. Compared with other informal caregivers, spousal caregivers are more likely to sole caregivers in the end of life of persons with disabilities [7], and most often in high-intensity caregiving [8]. Overall, spousal caregivers' negative experiences are dominant when caregiving, especially for who care for severely disabled older in China [9].

Disability inclusion is critical to achieving health for all, countries have an obligation to address the health inequities faced by persons with disabilities. In 2016, the Chinese government launched a targeted disability inclusion action— Long-term care insurance (LTCI)—in 15 cities. LTCI is part of China's national health policy and systems research agenda on disability, it prioritizes health equity for persons with disabilities, and provides a continuum of care in the form of cash or person-centered basic life care services and basic medical care services at home or

in institutions, with approximately 70% reimbursement from the LTCI fund. In addition, LTCI includes family caregivers, especially women caregivers, in the paid labour force, and economically empowers them.

As recipients, persons with disabilities were directly affected by LTCI. Previous literature found that LTCI not only reduced instrumental activities of daily living (IADL) and activities of daily living (ADL) scores among middle-aged and older adults, with urban residents benefited more [10, 11], but also improved their self-rated health [12] and reduced their depressive symptoms [13–16]. Compared with healthy older adults, LTCI had a more statistically significant positive effect on depression, mental state and episodic memory in older persons with disabilities by reducing healthcare costs, increasing daily companionship and social interaction [17]. LTCI beneficiaries were 8.8% more likely to self-report better health and 2.72 days longer hospital stays than non-beneficiaries [18]. Previous literature also found that LTCI not only reduced outpatient visits, hospital stays, hospital costs, and health insurance expenditures by 8.1%, 41.0%, 17.7%, and 11.4%, respectively [19], but also benefited mortality, survival time, and ADL in older adults with disabilities, and the effect of LTCI on mortality reduction was more pronounced in areas with abundant care resources [20].

LTCI also affected caregivers' health outcomes, because it included family caregivers in the paid labour force, economically empowered them. Previous literature found that caregiver was a stressful role [21], with higher rates of depressive symptoms compared with non-caregivers [22], and higher role overload was associated with more depressive and lower psychological well-being among spousal caregivers of older adults with disabilities [23]. A cross-country study found that respite care and nursing allowance reduced the deterioration of self-rated health among family caregivers [24], because burden care on family caregivers was reduced after using formal care provided by LTCI [25], with probability and intensity of informal care use were reduced by 5.7% and 17.4%, respectively, but there was no statistically significant policy effect for older adults with high or low incomes [26]. Previous literature also found that LTCI not only reduced the burden on informal caregivers, but also increased their participation in the labor market, there was a more pronounced reduction in care burden among spouses and LTCI primarily benefited informal carers who provided care for low-income or farmer older people [27]. Compared with urban spousal caregivers in China, rural spousal caregivers benefited more from LTCI on health [28]. However, study from Germany also found that home care allowance provided by LTCI did not affect caregivers' physical health [29].

The effect of LTCI also spilled over to persons who without disabilities and are non-caregivers. Previous

Jiang and Yang BMC Public Health (2024) 24:2262 Page 3 of 12

literature found that LTCI had beneficial health effect on non-caregiver family members, with greater benefits for family members with lower education levels and lower household incomes [30]. LTCI improved self-rated health among older adults, because LTCI might has a reassuring effect [31], while for middle-aged and older adults, LTCI was only effective for urban residents rather than rural residents [32]. LTCI also statistically significant reduced the depression in middle-aged and older adults [11, 33], as well as out-of-pocket inpatient and outpatient costs [12], number of outpatient visits, hospitalizations and days in hospital [11, 34]. However, Previous studies also found that the effect of LTCI on depression might not be significant, because improvement in depression might take a longer time [12], and there was no statistically significant reduction in the number of chronic diseases due to LTCI [35].

Efforts in achieving health for all must focus on reaching people most often left behind such as marginalized, stigmatized and geographically isolated people, with a particular focus on those in situations of increased vulnerability. Compared with men, women are disadvantaged by discrimination rooted in socio-cultural factors in many societies, such as unequal power relationships, and social norms associated with women that decrease education and paid employment opportunities, resulting the health outcomes of women are of particular concern. Compared with others, persons with disabilities and caregivers are also in situations of increased vulnerability. Women's health vulnerability is evident compared with men, however, the health vulnerability in women with different roles is inconsistent and unclear. Therefore, the objectives of our study were evaluating the effect of LTCI on health outcomes for women in different roles (rather than women compared with men), including female recipients (i.e., women with disabilities), female caregivers, and female non-recipients and female non-caregivers, using staggered difference-in-differences (DID) method and nationally representative health survey data, then discussing the heterogeneity of this effect based on geographic regions and urban-rural disparities.

Our contributions were reflected in the following aspects. Firstly, compared with the current literature, we looked beyond the single perspective of recipients or caregivers to women groups, who were previously neglected with a higher prevalence of disability, the majority of global care workforce and poorer health outcomes, and evaluated the effect of LTCI on women health from a comprehensive perspective of recipients, caregivers, and non-recipients and non-caregivers. Secondly, in practical contributions, our study demonstrated that the effect of LTCI on women's health varied across roles, geographic regions, and urban-rural disparities, which was conducive to deliver differentiated health interventions

for recipients, caregivers, and non-recipients and non-caregivers in the implementation process of LTCI in developing countries, with a focus on marginalized, stigmatized, and geographically isolated groups of women. Thirdly, in social contributions, our study demonstrated that LTCI achieved tripartite welfare improvements for female recipients, female caregivers, and female non-recipients and female non-caregivers, which provides policy implications for reaching people most often left behind and in situations of vulnerability, as well as minimizing health inequalities within women.

Methods

Data

Our data drew from China Health and Retirement Longitudinal Study (CHARLS). CHARLS collected a high quality nationally representative sample of Chinese residents aged 45 and over, adopting multi-stage stratified PPS sampling. CHARLS questionnaire included basic personal and community information, family structure and financial support, health status and physical measurement, medical service utilization and work, and retirement and pension. The baseline wave was fielded in 2011, including 10,000 households and 17,500 individuals in 150 counties/districts and 450 villages/resident committees. These samples were tracked every two or three years thereafter, currently 2020 wave was the latest update. However, previous information had many missing values in 2020 wave, due to the addition of coronavirus disease pandemic module and the simplification of information in other modules. Therefore, we retained data from 2011, 2013, 2015 and 2018 wave in CHARLS.

Regarding sample screening, we preserved samples of women aged 45 and over and removed missing values for related variables. Recipients in our study were identified as persons with disabilities who had difficulties with ADL or IADL. ADL reflected the number of items that respondents have difficulties in six basic activities, such as bathing, dressing, eating, getting in and out of bed, going to the toilet, and controlling urine, value ranged from 0 to 6, the higher the scores, the worse the ADL. IADL reflected the number of items in which respondents have difficulties in five instrumental activities, such as financial management, taking medicine, shopping, cooking, and housework, value of IADL ranged from 0 to 5, the higher the scores, the worse the IADL. Caregivers in our study referred to spousal caregivers, who helped their spouse (not parents, offspring, or other people) with any ADL or IADL. Non-recipients and non- caregivers were identified as a collection of people neither recipients nor spousal caregivers. To eliminate interference, we excluded cities that implemented LTCI on their own without the approval of central government. Finally, we obtained 16,707 samples of women aged 45 and over, including

3,962 in 2011 wave, 3,833 in 2013 wave, 4,064 in 2015 wave and 4,848 in 2018 wave.

Measures

The dependent variables were health outcomes. Health outcomes were measured with representative indicators: self-rated health [12, 31, 36, 37], depression [13–15, 28], and chronic diseases [35]. Self-rated health corresponded to the question in CHARLS "How would you rate your health status?". The response options were reversed encode to "1 very poor, 2 poor, 3 fair, 4 good, and 5 very good", then we standardized the response options by converting continuous variables into a dummy variable, where "1 very poor, 2 poor" was encoded as "0 poor", otherwise encoded as "1 good". Depression levels were calculated from the Center for Epidemiologic Studies Depression Scale-10 (CESD-10) with value ranged from 0 to 30, a score of 10 and above was considered depressed, and the higher the scores, the more severe the depression. Chronic diseases referred to the number of chronic diseases that respondents suffered from.

The independent variable was the implementation of LTCI in pilot cities, which involved two key data— Treat and Post. Treat represented cities (pilot and non-pilot cities), Post represented policy periods (before and after LTCI). Value range of LTCI was 0 to 1. Value 1 indicated that this city not only belonged to national LTCI pilot cities, but also posted LTCI in survey year. Value 0 indicated that this city didn't belong to national LTCI pilot cities or didn't post LTCI in survey year. Figure 1 reports the evolution of 15 national LTCI pilot cities. Only a few cities implemented LTCI in 2016 and before, and most cities did so in 2017 and later.

All analyses included series of control variables associated with women's health outcomes according to previous studies, particularly social determinants of health,

risk factors, and health system factor [17, 28, 32, 33, 38–42]. Control variables were specifically measured by age, public pension, education, employment, living conditions, social participation, residency arrangements, number of children, marriage status, smoking, alcohol consumption, future ADL help, and social medical insurance.

Table 1 reports definition and descriptive statistics of main variables in 16,707 women. Overall, 70.92% of women self-rated their health as good, their average depression levels exceeded 9.29, and 1.78 was their average number of chronic diseases. About 6.21% of women lived in cities where had implemented LTCI policy. Their average age was 60.77 years old, 33.99% of women received public pension, and only 22.15% of women had an education level of middle school and above. 56.77% of women worked in any job in the past year, 61.61% of women lived in rural village, and 48.09% of them participated in any social activities within the past month. 52.71% of women co-lived with their children, 2.82 was their average number of children, and only 0.07% of women unmarried. 5.36% of women currently had the habit of smoking and 13.81% of women had alcohol consumption within the past 12 months. Only 1.26% of women felt that professionals will be able to help them with their ADL needs in the future, while 93% of women participated in social medical insurance.

Statistical analysis

The staggered difference-in-differences (DID) method is a quasi-experimental technique for constructing a counterfactual framework [43]. Due to the staggered treatment timing, we established the staggered DID method with reference to other scholars [44]:

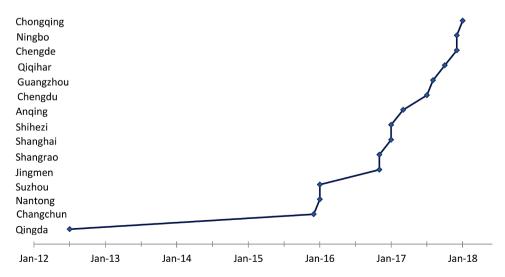


Fig. 1 Evolution of 15 national LTCI pilot cities

Jiang and Yang BMC Public Health (2024) 24:2262 Page 5 of 12

Table 1 Definition and descriptive statistics (N = 16,707)

Variables	Definition	Mean	Std. dev	Min	Max
Self-rated health	Self-rated good health = 1, self-rated poor health = 0	0.7092	0.4542	0	1
Depression	CES-D-10 score, ranged from 0 to 30	9.2972	6.6316	0	30
Chronic diseases	The number of chronic diseases that respondents suffered from	1.7866	1.5856	0	10
Treat*Post	City was pilot city, treat = 1, otherwise treat = 0. City posted LTCI in survey year, post = 1, otherwise post = 0	0.0621	0.2414	0	1
Age	Age of respondent in the survey year	60.7735	10.0059	45	100
Public pension	Received public pension, yes = 1, no = 0	0.3399	0.4737	0	1
Education	Middle school and above = 1, elemen- tary school and below = 0	0.2215	0.4152	0	1
Employment	Worked in any job in the past year, yes=1, no=0	0.5677	0.4954	0	1
Living conditions	Lives in rural village=1, lives in urban community=0	0.6161	0.4864	0	1
Social participation	Participated in any social activities within the past month, yes=1, no=0	0.4809	0.4996	0	1
Residency arrangements	Co-living with children, yes = 1, $no = 0$	0.5271	0.4993	0	1
Children	Number of children	2.8228	1.4411	1	10
Marriage	Unmarried = 1, married = 0	0.0007	0.0268	0	1
Smoking	Currently smoking, $yes = 1$, $no = 0$	0.0536	0.2253	0	1
Alcohol consumption	Alcohol consumption within the past 12 months, yes = 1, no = 0	0.1381	0.3451	0	1
Future ADL help	Professionals will be able to help them with their ADL needs in the future, yes=1, no=0	0.0126	0.1114	0	1
Social medical insurance	Participated in social medical insurance, yes=1, no=0	0.9300	0.2551	0	1

$$HEA_{ict} = \alpha_1 + \theta_1 Treat_{ic} * Post_{ct} + \lambda_1 Z_{ict} + \eta_c + \mu_t + \epsilon_{ict}$$
(1)

Where ${\rm HEA_{ict}}$ denotes health outcomes of women i who live in city c in time t. ${\rm Treat_{ic}}$ represents the treated group status (i.e., pilot list status) of women i who live in city c. ${\rm Post_{ct}}$ represents LTCI post status of city c in time t. θ_1 measures the effect of LTCI on health outcomes in women aged 45 and over. λ_1 is a vector of control variables $Z_{\rm ict}.$ η_c and μ_t represent city and year fixed effect. $\epsilon_{\rm ict}$ represents random perturbations that affect health. Finally, the standard errors were clustered at city level to correct for possible autocorrelation and heteroscedasticity.

The parallel trend hypothesis is a key prerequisite for constructing staggered DID method, which requires women's health outcomes trends in pilot and non-pilot cities must be parallel before implementation of LTCI. Therefore, using the event-study method proposed by other scholars [45], we established a parallel trend test model:

$$HEA_{ict} = \alpha_1 + \theta_t \sum_{-3}^{3} Treat_{ic} * Post_{ct}$$

$$+ \lambda_1 Z_{ict} + \eta_c + \mu_t + \epsilon_{ict}$$
(2)

Where $\theta_{\rm t}$ reflects the health outcomes disparities in pilot and non-pilot cities in time t of LTCI policy posted. There were few data 4 years before LTCI and 3 years after LTCI, so we aggregated data 4 years before LTCI into year -3, data 3 years after LTCI into year 3, and considered year -4 as the base year. Other variables are synonymous with Eq. (1).

Results

Main results

Table 2 reports the results obtained by staggered DID model. In women samples, the coefficient of Treat*Post was 0.0289 (significant level was 5%) in self-rated health, -0.7727 (significant level was 1%) in depression, and -0.0447 in chronic diseases (not significant). The coefficient of Treat*Post was 0.0664 (significant level was 5%) in self-rated health in the sample of female recipients, -0.3767 in chronic diseases (significant level was 10%) in the sample of female caregivers, -0.7754 (significant level was 1%) in depression in the sample of female non-recipients and female non-caregivers, and the coefficient of Treat*Post was not significant in other columns.

The results indicated that compared with women in non-pilot cities, LTCI statistically significant increased self-rated health and reduced depression levels in women in pilot cities, and improved the health in women with different roles by increasing self-rated health in female recipients, reducing the number of chronic diseases in

 Table 2
 The results of staggered DID model in women with different roles

Marine	Western			1000	24.00		2000			2000	A Language of the Contract of	1
Valiables				remaie recipients	cipiello		reillale calegivels	regivei s		non-caregivers	remale nom-recipiems and lemale non-caregivers	ש
	Self-rated health	Depression	Chronic diseases	Self- rated health	Depression	Chronic diseases	Self- rated health	Depression	Chronic diseases	Self-rated health	Depression	Chronic diseases
Treat*Post	0.0289**	-0.7727***	-0.0447	0.0664**	-0.7674	0.0094	9890:0	-0.3589	-0.3767*	0.0071	-0.7754***	-0.012
	(-0.0135)	(-0.1884)	(-0.0404)	(-0.0322)	(-0.6817)	(-0.1094)	(-0.0528)	(-0.6519)	(-0.1991)	(-0.0164)	(-0.2375)	(-0.0461)
Age	-0.0011*	0.0131	0.0203***	0.0039***	-0.0602***	0.0026	0.0026	-0.0029	0.0127	-0.0003	-0.0105	0.0214***
	(-0.0006)	(-0.0104)	(-0.0027)	(-0.001)	(-0.0149)	(-0.0036)	(-0.0023)	(-0.032)	(-0.0087)	(-0.0007)	(-0.0109)	(-0.0029)
Public pension	-0.0189**	-0.0098	0.0736**	-0.0084	-0.1427	0.0721	0.0013	-0.5307	-0.0347	-0.0251**	0.039	0.044
	(-0.0091)	(-0.142)	(-0.0333)	(-0.0168)	(-0.2184)	(-0.0571)	(-0.0338)	(-0.4153)	(-0.1145)	(-0.0101)	(-0.1741)	(-0.0347)
Education	0.0426***	-1.0718***	0.0085	0.0258	-1.2173***	0.2468***	-0.0372	0.0788	0.0251	0.0276**	-0.6784***	0.0363
	(-0.0113)	(-0.1761)	(-0.0528)	(-0.0203)	(-0.3328)	(-0.0938)	(-0.0465)	(-0.7427)	(-0.1365)	(-0.0113)	(-0.15)	(-0.0523)
Employment	0.1075***	-0.5972***	-0.3852***	0.1367***	-0.7032***	-0.3181***	0.1306***	-0.4107	-0.4395***	0.0520***	0.0202	-0.3237***
	(-0.0085)	(-0.1086)	(-0.0327)	(-0.0145)	(-0.2079)	(-0.0604)	(-0.0307)	(-0.4373)	(-0.1043)	(-0.0092)	(-0.1279)	(-0.0362)
Living	-0.0748***	1.0715***	0.0045	-0.0443*	0.2654	-0.2394**	-0.0992**	1.8782**	-0.0355	-0.0554***	0.9075***	0.0412
conditions												
	(-0.015)	(-0.2328)	(-0.0541)	(-0.0259)	(-0.2702)	(-0.0952)	(-0.0413)	(-0.7254)	(-0.2196)	(-0.0134)	(-0.2278)	(-0.0481)
Social	0.0449***	-0.8013***	0.032	0.0587***	-0.9934***	0.1163**	0.0515**	-1.4748***	0.0155	0.0149*	-0.2527*	0.0475
participation												
	(-0.0087)	(-0.123)	(-0.026)	(-0.016)	(-0.2083)	(-0.0487)	(-0.0234)	(-0.3322)	(-0.069)	(-0.0082)	(-0.132)	(-0.031)
Residency	-0.0068	0.0358	-0.0819**	-0.0004	-0.2979	-0.1493**	0.0234	-0.2976	-0.1992*	0.0001	0.0555	-0.0747**
arrangements												
	(-0.0088)	(-0.1329)	(-0.0324)	(-0.0156)	(-0.2373)	(-0.0575)	(-0.0316)	(-0.429)	(-0.1131)	(-0.0091)	(-0.1271)	(-0.0331)
Children	-0.0095***	0.0734	0.016	-0.0115**	0.0331	0.0147	-0.0093	0.1436	0.0166	-0.0047	0.0496	0.0109
	(-0.0035)	(-0.0609)	(-0.0159)	(-0.005)	(-0.0855)	(-0.0221)	(-0.0103)	(-0.1728)	(-0.0436)	(-0.0044)	(-0.0677)	(-0.0176)
Marriage	0.0152	-1.2895***	-0.6637***	0.0353***	-0.9242***	***/699'0-	1	ı	ı	-0.025	-1.8029***	-0.5524***
	(-0.039)	(-0.2159)	(-0.0704)	(-0.0095)	(-0.1247)	(-0.0337)	ı	1	1	(-0.0765)	(-0.1936)	(-0.1931)
Smoking	-0.0229	1.0369***	0.0926	-0.0193	0.9216**	0.0481	0.0129	0.7015	-0.2048	-0.0397*	1.1354***	0.1285
	(-0.0206)	(-0.3309)	(-0.0915)	(-0.0318)	(-0.4235)	(-0.1117)	(-0.0458)	(-0.8582)	(-0.1996)	(-0.0224)	(-0.3799)	(-0.0883)
Alcohol	0.0503***	0.1098	-0.2547***	0.0765***	-0.0755	-0.3436***	0.1047***	0.3604	-0.2262*	0.0398***	9660.0	-0.2128***
consumption												
	(-0.0108)	(-0.1674)	(-0.0439)	(-0.0231)	(-0.2591)	(-0.0784)	(-0.0349)	(-0.5912)	(-0.1151)	(-0.0101)	(-0.1739)	(-0.0439)
Future ADL help	0.0392	-1.3904***	-0.2359***	0.0442	-0.9406	-0.1475	0.1258	-4.2538***	-0.6967	0.0246	-1.1988***	-0.1828*
	(-0.0286)	(-0.3436)	(-0.0882)	(-0.0639)	(-0.7843)	(-0.1789)	(-0.1369)	(-1.4769)	(-0.5722)	(-0.0293)	(-0.3895)	(-0.1018)
Social medical insurance	-0.0037	-0.0285	0.1857***	-0.0101	0.1591	0.1173	-0.0718	0.0977	0.3052**	-0.0107	0.0107	0.2474***
	(-0.0132)	(-0.2063)	(-0.0447)	(-0.021)	(-0.3403)	(-0.0774)	(-0.044)	(-0.7537)	(-0.1525)	(-0.0155)	(-0.2521)	(-0.0544)
Fixed city	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Fixed year	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
_cons	0.7619***	8.6019***	0.5856***	0.2397***	16.9456***	2.3891***	0.5141***	10.9574***	1.4270**	0.8393***	7.7680***	0.1375

0.141

Chronic diseases

Depression

Self-rated

health

Chronic diseases

Depression

Self-

Female caregivers

Female non-recipients and female

non-caregivers

(-0.1722) 10,108

(-0.6155) 10,108 0.086

(-0.0423)

(-0.5935)

(-2.2281)

rated health (-0.1397) 10,108

1631

1631

1631

diseases Chronic (-0.2426)5747 Depression (-1.0319)Female recipients 5747 (-0.0767)health rated 5747 Selfdiseases Chronic (-0.1535)16,707 Depression (-0.6491)16,707 Self-rated Women health (-0.039)16,707 /ariables

Table 2 (continued)

0.238 Vote *, **, and *** represent significant levels of 10%, 5%, and 1%, respectively. Values in parentheses are city-level clustering robust standard errors 0.163 0.118 0.188 0.096 0.078 0.166 0.106 0.078

female caregivers, and reducing depression levels in female non-recipients and female non-caregivers. However, there was no evidence that LTCI statistically significant improved the health among women with different roles in pilot cities in other situations.

Parallel trend and robustness tests

Table 3 reports the parallel trend test results obtained by event-study method at 95% confidence interval with considering year -4 as the base year. The results indicated that overall the coefficient of Treat*Post were not statistically significant in year -3, year -2 and year -1 (except for self-rated health among women in year -3 and among female recipients in year -1), while in year 0 and after, they were mostly not only statistically significant, but also preserved correct treatment effect signs. Therefore, our sample generally passed parallel trend test.

To evaluate the potential bias risks associated with staggered DID and increase the credibility in results [46], we conducted robustness checks from three aspects. Firstly, considering that even in the absence of time-varying processing points, some covariates may lead to bias in the two-way fixed effects estimates, we reported results of the stacked regression without covariates in women samples to understand the robustness of the effect estimates and the degree to which they rely on the inclusion of controls. Secondly, we redefined dependent variables in samples of women with different roles. Specifically, self-rated health was redefined by continuous variables, chronic diseases number was redefined by prevalence of chronic disease comorbidities, and depression levels was redefined by prevalence of depression. Thirdly, we conducted a placebo test by randomly selecting the treated group and pilot time and iterating this random selection process 500 times.

Table 4 reports the results of robustness test 1 and test 2. The results indicated that LTCI still statistically significant increased women's self-rated health and reduced depression, increased self-rated health for female recipients, reduced chronic diseases for female caregivers, and reduced depression for female non-recipients and female non-caregivers, which were consistent with those in staggered DID model.

Figure 2 reports kernel density distribution of the regression coefficients across 500 iterations of simulations for women, female recipients, female caregivers, and female non-recipients and female non-caregivers. It can be found that the regression coefficients were centrally distributed around the 0 value, obeying the normal distribution, while the actual estimated coefficients (vertical dotted line) were obviously an outlier or far away from the 0 value, which indicated that the randomness factor had no statistically significant effect on our results,

Jiang and Yang BMC Public Health (2024) 24:2262 Page 8 of 12

Table 3 Parallel trend test

Variables	Women	Women	Female recipients	Female caregivers	Female non-recipients and female non-caregivers
	Self-rated health	Depression	Self-rated health	Chronic diseases	Depression
Treat*Post year – 3	0.0443*	-0.2056	0.0297	-0.1441	0.0016
	(0.0254)	(0.3977)	(-0.0632)	(-0.2404)	(-0.5643)
Treat*Post year – 2	0.0016	-0.4991	0.0335	-0.0308	-0.3761
	(0.0156)	(0.3662)	(-0.0573)	(-0.2842)	(-0.3964)
Treat*Post year – 1	-0.0304	0.3621	-0.1464*	-0.1742	-0.1981
	(0.0421)	(0.5799)	(-0.0815)	(-0.2398)	(-0.8994)
Treat*Post year 0	0.0488**	-1.0684***	0.1207**	-0.4981	-0.6642
	(0.0204)	(0.3475)	(-0.0483)	(-0.3609)	(-0.4384)
Treat*Post year 1	0.0293	-0.9513***	0.0619	-0.5246**	-0.9928 ^{**}
	(0.0275)	(0.3433)	(-0.0615)	(-0.2363)	(-0.4135)
Treat*Post year 2	0.0241	-0.4778*	-0.057	-0.7035*	-1.1418**
	(0.0233)	(0.2518)	(-0.0657)	(-0.3985)	(-0.5659)
Treat*Post year 3	0.0098	-0.9494**	0.0326	0.0475	-0.3205
	(0.0196)	(0.4712)	(-0.0325)	(-0.2167)	(-0.5231)
Control variables	Yes	Yes	Yes	Yes	Yes
Fixed city	Yes	Yes	Yes	Yes	Yes
Fixed year	Yes	Yes	Yes	Yes	Yes
_cons	0.7606***	8.6210***	0.2397***	1.4468**	7.7797***
	(0.0389)	(0.6521)	(-0.0768)	(-0.599)	(-0.6196)
N	16,707	16,707	5747	1631	10,108
R^2	0.078	0.106	0.079	0.239	0.086

Note *, **, and *** represent significant levels of 10%, 5%, and 1%, respectively. Values in parentheses are city-level clustering robust standard errors

Table 4 Robustness test 1 and test 2

Variables	Without covariates		Redefine dependen	t variables	
	Women		Female recipients	Female caregivers	Female non-recipients and female non-caregivers
	Self-rated health	Depression	Self-rated health	Chronic diseases	Depression
Treat*Post	0.0257*	-0.7332***	0.1607***	-0.1296*	-0.0493***
	(0.0136)	(0.1876)	(0.0554)	(0.0703)	(0.0186)
Control variables	No	No	Yes	Yes	Yes
Fixed city	Yes	Yes	Yes	Yes	Yes
Fixed year	Yes	Yes	Yes	Yes	Yes
_cons	0.7076***	9.3427***	1.8659***	0.5093***	0.3137***
	(0.0008)	(0.0117)	(0.1345)	(0.1496)	(0.0490)
N	16,707	16,707	5743	1631	10,108
R^2	0.047	0.086	0.089	0.207	0.066

Note *, **, and *** represent significant levels of 10%, 5%, and 1%, respectively. Values in parentheses are city-level clustering robust standard errors

the placebo test passed. Therefore, our results were generally robust.

Analysis of heterogeneity

We explored the geographical and social heterogeneity of LTCI's effect on women's health outcomes. Table 5 reports the geographical heterogeneity results for women. The results indicated that in the west and northeast of China, compared with women in non-pilot cities, LTCI statistically significant reduced women's depression (the coefficient of Treat*Post was -0.6324, significant level was 5%) and chronic diseases (the coefficient

of Treat*Post was -0.1544, significant level was 5%) in pilot cities. In the east and central, LTCI statistically significant reduced women's depression (the coefficient of Treat*Post was -0.7628, significant level was 1%) in pilot cities compared with women in non-pilot cities. In other words, there was a more pronounced improvement in health outcomes among women in the west and northeast.

Table 6 reports the social heterogeneity results of women. The results indicated that in rural village, compared with women in non-pilot cities, LTCI statistically significant increased women's self-rated health (the

Jiang and Yang BMC Public Health (2024) 24:2262 Page 9 of 12

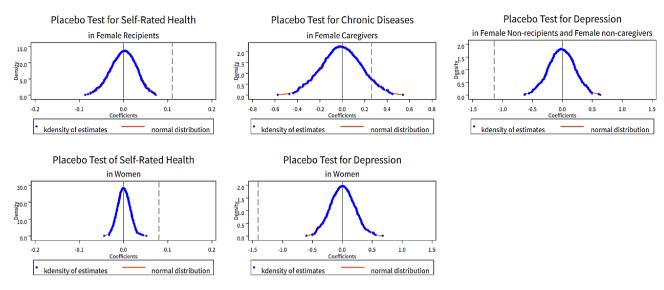


Fig. 2 Placebo test in women with different roles

Table 5 Geographical heterogeneity of women

Variables	Women in the west	and northeast		Women in the east	and central	
	Self-rated health	Depression	Chronic diseases	Self-rated health	Depression	Chronic diseases
Treat*Post	0.0521	-0.6324**	-0.1544**	0.0163	-0.7628***	0.0024
	(-0.036)	(-0.3079)	(-0.0765)	(0.0141)	(0.2355)	(0.0503)
Control variables	Yes	Yes	Yes	Yes	Yes	Yes
Fixed city	Yes	Yes	Yes	Yes	Yes	Yes
Fixed year	Yes	Yes	Yes	Yes	Yes	Yes
_cons	0.7298***	9.2846***	0.7716***	0.7838***	8.1230***	0.4523**
	(-0.0537)	(-1.1518)	(-0.2155)	(-0.0572)	(-0.7211)	(-0.2113)
N	7371	7371	7371	9336	9336	9336
R^2	0.076	0.104	0.158	0.076	0.098	0.165

Note *, **, and *** represent significant levels of 10%, 5%, and 1%, respectively. Values in parentheses are city-level clustering robust standard errors

Table 6 Social heterogeneity of women

Variables	Women live in rural	village		Women live in urba	n community	
	Self-rated health	Depression	Chronic diseases	Self-rated health	Depression	Chronic diseases
Treat*Post	0.0464**	-1.2605***	-0.0339	0.0095	-0.2042	-0.0523
	(0.0184)	(0.2140)	(0.0609)	(0.0195)	(0.2594)	(0.0634)
Control variables	Yes	Yes	Yes	Yes	Yes	Yes
Fixed city	Yes	Yes	Yes	Yes	Yes	Yes
Fixed year	Yes	Yes	Yes	Yes	Yes	Yes
_cons	0.7321***	9.8708***	1.1016***	0.7055***	8.1527***	-0.1099
	(0.0483)	(0.8361)	(0.1812)	(0.0562)	(0.7929)	(0.2384)
N	10,293	10,293	10,293	6414	6414	6414
R^2	0.080	0.092	0.167	0.078	0.112	0.198

Note *, **, and *** represent significant levels of 10%, 5%, and 1%, respectively. Values in parentheses are city-level clustering robust standard errors

coefficient of Treat*Post was 0.0464, significant level was 5%) and reduced the depression (the coefficient of Treat*Post was -1.2605, significant level was 1%) in pilot cities. However, there was no evidence that LTCI statistically significant affected health outcomes for women lived in urban community in pilot cities compared with women in non-pilot cities. In other words, there was

a more pronounced improvement in health outcomes among women in rural village.

Discussion

Using nationally representative sample and staggered DID method, we demonstrated that compared with women in non-pilot cities, LTCI statistically significant increased self-rated health and reduced depression in

Jiang and Yang BMC Public Health (2024) 24:2262 Page 10 of 12

women in pilot cities, and improved the health in women with different female roles. Specifically, compared with female counterparts in non-pilot cities, LTCI statistically significant increased female recipients' self-rated good health, reduced female caregivers' chronic diseases, and reduced female non-recipients' and female non-caregivers' depression in pilot cities.

Possible explanations were as follows. On the one hand, for recipients, who were more vulnerable than persons without disabilities, the implementation of LTCI optimizing their affordability, accessibility, and quality of care services, alleviating their disability levels [10, 11], reducing medical costs [19], and increasing daily companionship and social interaction they obtained [17], thereby improving the health outcomes of female recipients. Our results were consistent with previous studies [12–16]. On the other hand, for caregivers, caregiver was a stressful role [21], spousal caregivers' negative experiences were dominant while caregiving, with higher financial and physical burden, depressive symptoms [6] and highintensity caregiving [8]. However, the implementation of LTCI reduced their caregiving burden [25], especially for spousal caregiver [27, 30], reduced their probability and intensity of informal care use [26], thereby improving the health outcomes of female caregivers. For non-disabled women without providing care, even though they did not obtain care services and cash grants, LTCI make them reassured by providing protection against the risk of future disability, and reduced their probability of providing care and increased the duration of nighttime sleep [30], thereby improving the health outcomes of female non-recipients and female non-caregivers. Our results were supported by previous studies [12, 18, 24, 28, 33].

We also demonstrated that the effect of LTCI on women's health outcomes was geographically and socially heterogeneous. Compared with women in in the east and central and women in urban community, the positive effect of LTCI on women's health outcomes was more pronounced in women in the west and northeast and women in rural village. Potential explanations were as follows. Firstly, we consider the reason why the positive effect of LTCI on women's health outcomes was more pronounced in women in the west and northeast was that regional economic differences produced differences in burden of care on family. Compared with the east and central where the economic level was better in China, the west and northeast where the economic level was worse, the purchasing power for formal care services was limited, resulting in a heavier burden of care on families. After the implementation of LTCI, families with disabilities in the west and northeast could obtained care services or cash benefits from LTCI, which greatly alleviated their burden of care.

Additionally, we consider the reason why the positive effect of LTCI on women's health outcomes was more pronounced in women in rural village was that the ruralurban gap in chronic investment in health and care work. Chronic underinvestment in rural village was worse than that in urban community, women in rural village may undertake more unpaid care work and burden of care on family was heavier. It was estimated that 51-67% of rural population couldn't obtain adequate basic health services, and in some countries the number of health workers available among rural population was 10 times lower than the number among urban population [47], the rate of medical rehabilitation services utilization in urban areas was almost twice that of rural village [48]. After the impletion of LTCI, the burden of care on family in rural village was greatly alleviated and caregivers were included more equitably in the paid labour workforce by obtaining care services or cash benefits.

The limitations of our study were that data on women's health outcomes in CHARLS was self-rated and may be susceptible to memory biases. In addition, factors that influence women's health outcomes were abundant, while we just controlled some of them. These limitations could be improved in future studies.

Conclusions

In summary, our study found that after the implementation of LTCI in China, health outcomes in women, including female recipients, female caregivers, and female non-recipients and female non-caregivers, were statistically significant improved, and the effect of LTCI on women's health outcomes was geographically and socially heterogeneous. Our findings highlight the importance of delivering differentiated health interventions for recipients, caregivers, and non-recipients and non-caregivers in the implementation process of LTCI, and minimizing health inequalities in geography and society within women.

Acknowledgements

Not applicable

Author contributions

Conceptualization, W.J.; methodology, H.Y. and W.J.; software, W.J; validation, H.Y. and W.J.; formal analysis, H.Y. and W.J.; investigation, H.Y. and W.J.; resources, H.Y.; data curation, W.J.; writing—original draft preparation, W.J.; writing—review and editing, H.Y. and W.J.; visualization, W.J.; supervision, H.Y. and W.J.; project administration, H.Y.; funding acquisition, H.Y. All authors read and approved the final manuscript.

Funding

This work is supported by "Health inequality related to air pollution: level measurement, policy evaluation and mechanism testing", the General Program of the National Natural Science Foundation of China (72274145); "Research on the effect evaluation and optimization path of long-term care insurance pilot policy based on counterfactual analysis", the Planning Fund Project of the Ministry of Education of the People's Republic of China (22YJA630101); "Research on the effect evaluation and optimization path of long-term care insurance pilot policy based on counterfactual analysis",

the Fundamental Research Funds for the Central Universities (413000138); "Research on the policy effect of long term care insurance", the self-set project of the Key Research Institute of Humanities and Social Sciences of the Ministry of Education of the People's Republic of China (21JDZS04).

Data availability

CHARLS data are publicly available at http://charls.pku.edu.cn/en/.

Declarations

Ethics approval and consent to participate

Not applicable.

Consent for publication

Not applicable.

Competing interests

The authors declare no competing interests.

Received: 12 April 2024 / Accepted: 14 August 2024

Published online: 20 August 2024

References

- WHO. Global report on health equity for persons with disabilities. 2022, 312, doi:https://www.who.int/publications/i/item/9789240063600
- ILO; WHO. The gender pay gap in the health and care sector 2022, doi:https:// www.who.int/publications/i/item/9789240052895
- WHO. Fair share for health and care: gender and the undervaluation of health and care work. 2024, doi:https://www.who.int/publications/i/ item/9789240082854
- Miyawaki CE, Chen NW, Meyer OL, Tran MT, Markides KS. Vietnamese adult-child and spousal caregivers of older adults in Houston, Texas: results from the Vietnamese aging and care survey (VACS). J Gerontol Soc Work. 2020;63:5–18. https://doi.org/10.1080/01634372.2019.1707735.
- Li L, Wister AV, Mitchell B. Social isolation among spousal and adult-child caregivers: findings from the Canadian longitudinal study on aging. Journals Gerontology: Ser B. 2021;76:1415–29. https://doi.org/10.1093/geronb/ gbaa197.
- Pinquart M, Sörensen S, Spouses. Adult children, and Children-in-Law as caregivers of older adults: a Meta-Analytic comparison. Psychol Aging. 2011;26:1–14. https://doi.org/10.1037/a0021863.
- Ornstein KA, Wolff JL, Lund EB, Rahman OK, Kelley AS. Spousal caregivers are caregiving alone in the last years of life. Health Aff. 2019;38:964–72. https:// doi.org/10.1377/hlthaff.2019.00087.
- 8. NAC; AARP. Research report: Careiving in the U.S. 2020, doi:https://www.caregiving.org/research/caregiving-in-the-us/caregiving-in-the-us-2020/
- Yang L, Wei W, Wu Y, Zhu S, Zeng X, Wang R, Zhang M, Lin X, Zhou C. The experiences of caring for disabled older adults in long-term: a qualitative study from the perspective of spousal caregivers. Chronic Illn. 2023;19:848– 61. https://doi.org/10.1177/17423953221148972.
- Cao N, Shi T, Pan C. Does long-term care insurance reduce the disability among middle-aged and older adults? Evidence from China. BMC Public Health. 2023;23:1138. https://doi.org/10.1186/s12889-023-16057-0.
- Tang Y, Chen T, Zhao Y, Taghizadeh-Hesary F. The impact of the long-term Care Insurance on the medical expenses and Health Status in China. Front Public Health 2022, 10, 847822, https://doi.org/10.3389/fpubh.2022.847822
- Wang J, Guan J, Wang G. Impact of long-term care insurance on the health status of middle-aged and older adults. Health Econ. 2023;32:558–73. https:// doi.org/10.1002/hec.4634.
- Song H, Sun H. Association of unmet long-term care needs with depression and anxiety among older adults in urban and rural China. Geriatr Nurs. 2023;49:115–21. https://doi.org/10.1016/j.gerinurse.2022.11.017.
- Tseng Y-T, Chen IH, Lee P-H, Lin P-C. Effects of auricular acupressure on depression and anxiety in older adult residents of long-term care institutions: a randomized clinical trial. Geriatr Nurs. 2021;42:205–12. https://doi. org/10.1016/j.gerinurse.2020.08.003.
- Efendi F, Tonapa SI, Has EMM, Ho KHM. Effects of chair-based resistance band exercise on physical functioning, sleep quality, and depression of older adults

- in long-term care facilities: systematic review and meta-analysis. Int J Nurs Sci. 2023;10:72–81. https://doi.org/10.1016/j.ijnss.2022.12.002.
- Kono A, Kanaya Y, Fujita T, Tsumura C, Kondo T, Kushiyama K, Rubenstein LZ. Effects of a Preventive Home visit Program in Ambulatory Frail Older people: a Randomized Controlled Trial. Journals Gerontology: Ser A. 2011;67A:302–9. https://doi.org/10.1093/gerona/glr176.
- Wang L. The impact of long-term care insurance pilot on the mental health of older adults: quasi-experimental evidence from China. SSM - Popul Health. 2024;25:101632. https://doi.org/10.1016/j.ssmph.2024.101632.
- Kim H, Mitra S. The Economic and Health effects of Long-Term Care Insurance: New evidence from Korea. J Econ Ageing. 2022;23:100412. https://doi. org/10.1016/j.jeoa.2022.100412.
- Feng J, Wang Z, Yu Y. Does long-term care insurance reduce hospital utilization and medical expenditures? Evidence from China. Soc Sci Med. 2020;258:113081. https://doi.org/10.1016/j.socscimed.2020.113081.
- Zeng L, Zhong Y, Chen Y, Zhou M, Zhao S, Wu J, Dong B, Dou Q. Effect of long-term care insurance in a pilot city of China: Health benefits among 12,930 disabled older adults. Arch Gerontol Geriatr. 2024;121:105358. https:// doi.org/10.1016/j.archger.2024.105358.
- 21. Portier C. Heterogeneity in the mental health cost of caring for others. Soc Sci Med. 2023;326:115922. https://doi.org/10.1016/j.socscimed.2023.115922.
- Capistrant BD. Caregiving for older adults and the caregivers' Health: an epidemiologic review. Curr Epidemiol Rep. 2016;3:72–80. https://doi. org/10.1007/s40471-016-0064-x.
- 23. Nah S, Martire LM, Zhaoyang RP, Gratitude. Role overload, and Mental Health among Spousal caregivers of older adults. J Gerontol B Psychol Sci Soc Sci. 2022;77:295–9. https://doi.org/10.1093/qeronb/gbab086.
- Calvó-Perxas L, Vilalta-Franch J, Litwin H, Mira P, Garre-Olmo J. A longitudinal study on public policy and the health of in-house caregivers in Europe. Health Policy. 2021;125:436–41. https://doi.org/10.1016/j.healthpol.2021.02.001.
- Chen M-C, Kao C-W, Chiu Y-L, Lin T-Y, Tsai Y-T, Jian Y-TZ, Tzeng Y-M, Lin F-G, Hwang S-L, Li S-R et al. Effects of home-based long-term care services on caregiver health according to age. Health Qual Life Outcomes 2017, 15, 208, https://doi.org/10.1186/s12955-017-0786-6
- Ning J, Chen H, Li M. Does income matter for the policy effect of public longterm care insurance on informal care use in China? A quasi-experimental study. J Health Serv Res Policy. 2024;13558196241252394. https://doi. org/10.1177/13558196241252394.
- Pei X, Yang W, Xu M. Examining the Impact of Long-Term Care Insurance on the Care Burden and Labor Market participation of Informal carers: a quasiexperimental study in China. Journals Gerontology: Ser B. 2024;79:gbae023. https://doi.org/10.1093/geronb/gbae023.
- Jiang W, Yang H. Health spillover studies of long-term care insurance in China: evidence from spousal caregivers from disabled families. Int J Equity Health. 2023;22(191). https://doi.org/10.1186/s12939-023-02001-6.
- Schmitz H, Westphal M. Short- and medium-term effects of informal care provision on female caregivers' health. J Health Econ. 2015;42:174–85. https://doi.org/10.1016/j.jhealeco.2015.03.002.
- Yi Y, Xin J, Liu J, Wu J. Health effects of long-term care insurance on spouses of disabled people: a quasi-experimental study. BMC Geriatr. 2023;23:679. https://doi.org/10.1186/s12877-023-04344-9.
- 31. Lei X, Bai C, Hong J, Liu H. Long-term care insurance and the well-being of older adults and their families: evidence from China. Soc Sci Med. 2022;296:114745. https://doi.org/10.1016/j.socscimed.2022.114745.
- Liu P, Yang Y, Yang Y, Cheng J. Different impact on health outcomes of longterm care insurance between urban and rural older residents in China. Sci Rep. 2023;13:253. https://doi.org/10.1038/s41598-023-27576-6.
- Jiang W, Yang H. Effect of long-term care insurance policy on depression in non-disabled people: evidence from China. BMC Public Health. 2024;24:954. https://doi.org/10.1186/s12889-024-18375-3.
- Chen H, Ning J. The impacts of long-term care insurance on health care utilization and expenditure: evidence from China. Health Policy Plan. 2022;37:717–27. https://doi.org/10.1093/heapol/czac003.
- Ma C, Yu Q, Song Z, Chen H. Long-term Care Insurance, the Control of Medical Expenses and value-based Health Care. China Industrial Econ. 2019;42–59. https://doi.org/10.19581/j.cnki.ciejournal.2019.12.003.
- Pan Y, Pikhartova J, Bobak M, Pikhart H. Reliability and predictive validity of two scales of self-rated health in China: results from China Health and Retirement Longitudinal Study (CHARLS). BMC Public Health. 2022;22:1863. https:// doi.org/10.1186/s12889-022-14218-1.

- Abdulrahim S, El Asmar K. Is self-rated health a valid measure to use in social inequities and health research? Evidence from the PAPFAM women's data in six arab countries. Int J Equity Health. 2012;11:53. https://doi. org/10.1186/1475-9276-11-53.
- Li X, Ge T, Dong Q, Jiang Q. Social participation, psychological resilience and depression among widowed older adults in China. BMC Geriatr. 2023;23:454. https://doi.org/10.1186/s12877-023-04168-7.
- Xu M, Yang W. Depressive symptoms among older empty nesters in China: the moderating effects of social contact and contact with one's children. Aging Ment Health. 2023;27(54–69). https://doi.org/10.1080/13607863.2021.2 019190.
- Gu L, Yu M, Xu D, Wang Q, Wang W. Depression in Community-Dwelling older adults living alone in China: Association of Social Support Network and functional ability. Res Gerontol Nurs. 2020;13:82–90. https://doi. org/10.3928/19404921-20190930-03.
- Huang M, Liu K, Liang C, Wang Y, Guo Z. The relationship between living alone or not and depressive symptoms in older adults: a parallel mediation effect of sleep quality and anxiety. BMC Geriatr. 2023;23:506. https://doi. org/10.1186/s12877-023-04161-0.
- 42. Pilehvari A, You W, Lin X. Retirement's impact on health: what role does social network play? Eur J Ageing. 2023;20:14. https://doi.org/10.1007/s10433-023-00759-w.
- Angrist JD, Krueger AB. Chapter 23 Empirical Strategies in Labor Economics. Handbook of Labor Economics 1999, 3, 1277–1366, https://doi.org/10.1016/ \$1573-4463(99)03004-7

- Beck T, Levine R, Levkov A. Big bad banks? The winners and losers from Bank Deregulation in the United States. J Finance. 2010;65:1637–67. https://doi. org/10.1111/j.1540-6261.2010.01589.x.
- Jacobson LS, LaLonde RJ, Sullivan DG. Earnings losses of Displaced workers. Am Econ Rev. 1993;83:685–709. http://www.jstor.org/stable/2117574.
- Baker AC, Larcker DF, Wang CCY. How much should we trust staggered difference-in-differences estimates? J Financ Econ. 2022;144:370–95. https://doi.org/10.1016/j.jfineco.2022.01.004.
- 47. WHO. WHO guideline on health workforce development, attraction, recruitment and retention in rural and remote areas. 2021, doi:https://www.who.int/publications/i/item/9789240024229
- Szybalska A, Broczek K, Slusarczyk P, Kozdron E, Chudek J, Puzianowska-Kuznicka M, Kostka T, Skalska A, Mossakowska M. Utilization of medical rehabilitation services among older poles: results of the PolSenior study. Eur Geriatr Med. 2018;9:669–77. https://doi.org/10.1007/s41999-018-0077-8.

Publisher's note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.